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Title: Confirmatory Factor Analysis and Psychometric Properties of the Norwegian Version of the Repetitive Eating-Questionnaire [Rep(eat)-Q]: Further Evidence of Two Distinct Subtypes of Grazing Behaviour

(\*identifying information removed to comply with journal peer-review policy)

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### Abstract

**Objective:** The Repetitive Eating–Questionnaire [Rep(eat)-Q] is a 12-item self-report measure of compulsive and non-compulsive forms of grazing behaviour (i.e., eating modest amounts of food in a repetitive and unplanned manner). The aim was to validate the proposed two-factor model of the Rep(eat)-Q in a community sample. **Method:** A total of 190 university students (78% female) were administered the Rep(eat)-Q along with other measures of eating behaviour. Mean age was 22.6 ( $SD=4.2$ , 19-43) and mean body mass index (BMI;  $\text{kg}/\text{m}^2$ ) was 22.4 ( $SD = 2.9$ , 17-37). **Results:** Findings revealed good fit indexes for the 2-factor model by CFA, supporting the original solution. Internal consistency was excellent for the total score and two subscales (range 0.86 to 0.91). Age and BMI did not correlate significantly with the Rep(eat)-Q, but moderate to strong correlations ( $r_{sp} \geq 0.48$  to 0.61) were found between the Compulsive Grazing (GC) subscale and eating-related measures, whereas weaker correlations ( $r_{sp} \geq 0.37$  to 0.45) were found for Repetitive Eating (RE). **Conclusions:** This study confirmed the proposed factor structure of the Rep(eat)-Q in a community sample, offering support to the conceptual distinction between compulsive (marked by loss of control) and repetitive (non-compulsive) subtypes of grazing behaviour.

*Keywords:* Grazing, loss of control, compulsive, repetitive eating, psychometric properties, self-report questionnaire

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## Highlights

- The Rep(eat)-Q demonstrated reliability and validity as a measure of grazing behaviour in a community sample of young Norwegian adults.
- Confirmatory factor analysis confirmed the originally proposed two-factor model of compulsive and non-compulsive/repetitive eating subtypes of grazing behavior.
- This study provides evidence of distinct subtypes of grazing behavior distinguished by loss of control (LOC): compulsive grazing (marked by LOC) and repetitive eating (eating in a distracted or mindless way).

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82 Confirmatory Factor Analysis and Psychometric Properties of the Norwegian Version of the

83 Repetitive Eating-Questionnaire [Rep(eat)-Q]: Further Evidence of Two Distinct Subtypes of

84 Grazing Behaviour

85 The repetitive eating of small/modest amounts of food in an unplanned manner, often

86 referred to as “grazing”, has been characterized as a pattern of unstructured and unanticipated

87 eating outside of planned meals and snacks , and/or not in response to sensations of hunger or

88 satiety (Conceicao, Mitchell, Engel, et al., 2014; Conceição et al., 2017). Grazing behaviour

89 has been investigated most extensively within treatment-seeking populations of eating

90 disorders (EDs) and obesity (see meta-analysis by Heriseanu, Hay, Corbit, &amp; Touyz, 2017),

91 especially obese bariatric samples at pre- and post-operative assessment, yet knowledge

92 regarding the nature and frequency of grazing behaviour in the community remains sparse.

93 Additional data would further inform the construct of grazing behavior, and clarify its

94 associations with features of disordered eating and body mass index (BMI) in different

95 populations across the BMI spectrum.

96 At present, it remains unclear whether grazing represents a high-risk or maladaptive

97 eating behaviour warranting specific clinical attention, or whether grazing lies outside the

98 spectrum of disordered eating. Grazing which persists or emerges post-operatively has been,

99 albeit inconsistently, associated with poorer weight loss outcomes in bariatric surgery (Colles,

100 Dixon, &amp; O'Brien, 2008; Conceicao, Mitchell, Vaz, et al., 2014; Kofman, Lent, &amp;

101 Swencionis, 2010; Robinson et al., 2014). A few studies have also demonstrated links

102 between grazing and various aspects of quality of life and mental health (Micanti et al., 2017;

103 Nicolau et al., 2015). Associations with eating disorder (ED) pathology, such as binge eating,

104 have been found, albeit inconsistently, within treatment-seeking individuals with eating

105 disorders or obesity (Conceicao, Bastos, et al., 2014; Conceicao, Mitchell, Vaz, et al., 2014;

106 Goodpaster et al., 2016), with relatively few corresponding studies in non-clinical populations  
107 (Heriseanu et al., 2017). Collectively, findings tend to diverge according to whether the  
108 construct of loss of control (LOC) is reflected in the definition and assessment of grazing. In  
109 particular, little to no evidence exists for an association between ED pathology and  
110 conceptualizations of grazing which do not capture LOC (i.e., picking or nibbling) (Conceicao  
111 et al., 2013; Masheb, Roberto, & White, 2013; Reas, Wisting, Kapstad, & Lask, 2012), in line  
112 with prior conceptualizations delineating pathological versus non-pathological forms of  
113 grazing based upon accompanying LOC (Lane & Szabo, 2013). Recent findings in a bariatric  
114 sample suggest that grazing lies on a continuum of LOC eating within the spectrum of ED  
115 psychopathology, with non-compulsive grazing associated with the lowest degree of  
116 LOC/psychopathology, whereas compulsive grazing lies more similarly to subjective and  
117 objective binge eating (Conceicao et al., 2018), highlighting the importance of compulsive  
118 grazing behaviour as a clinically relevant eating behaviour.

119 Varying definitions and inconsistent measurement of grazing and related eating  
120 behaviour has posed major challenges to this literature (for a comprehensive review, see  
121 Conceicao, Mitchell, Engel, et al., 2014). One attempt to operationalize grazing based upon  
122 expert consensus recently led to the development of a structured clinical interview and  
123 companion self-report questionnaire, the Repetitive Eating Questionnaire [Rep(eat)-Q]  
124 (Conceição et al., 2017). Items provide a dimensional assessment of two theoretically distinct  
125 subtypes of grazing, namely compulsive grazing (i.e., marked by a loss of control or feeling  
126 unable to resist eating despite attempts) and repetitive eating, or non-compulsive grazing (i.e.,  
127 eating in a distracted, mindless fashion). The initial development study supported the utility  
128 and validity of the Rep(eat)-Q as a brief, 12-item measure of grazing behaviour (Conceição et  
129 al., 2017). An exploratory factor analysis revealed a two-factor model, providing evidence for  
130 theoretical subtypes of compulsive versus non-compulsive grazing, showing a good fit by a

131 confirmatory factor analysis in an obese bariatric sample. Tests of convergent validity showed  
132 that compulsive grazing had stronger associations with core ED features than repetitive or  
133 non-compulsive grazing, in line with a priori assumptions. Additional validation studies are  
134 warranted to replicate the utility of the Rep(eat)-Q as a brief, self-report measure of grazing  
135 behaviour. This is also in line with efforts to standardize the assessment of grazing to allow  
136 meaningful comparisons and interpretation of findings across studies (Conceicao, Mitchell,  
137 Engel, et al., 2014; Heriseanu et al., 2017).

138 To date, there are no studies which have replicated and validated the two-factor  
139 structure of the Rep(eat)-Q aside from the initial development study. This is despite clear  
140 relevance for improving the evidence base for this newly-developed measure, and the value of  
141 such data to ascertain the frequency and clarify associations of grazing behaviour with ED  
142 pathology in the community. Thus, the aim of the present study was to examine the  
143 psychometric properties and test the proposed two-factor model of the Rep(eat)-Q using  
144 confirmatory factor analysis in a community sample.

## 145 **Method**

### 146 Participants

147 A total of 190 university students (78% female) were administered the Rep(eat)-Q  
148 along with other measures of eating behaviour. Mean age was 22.6 ( $SD= 4.2$ , 19-43) and self-  
149 reported mean BMI was 22.4 ( $SD = 2.9$ , 17-37). The majority of participants ( $n = 147$ , 80.3%)  
150 were classified in the healthy weight range ( $BMI = 18.5-24.9$ ), while 6 (3.2%) participants  
151 were classified as underweight, 26 (14.2%) as overweight and 4 (2.2%) as obese.

### 152 Measures

153 The Repetitive Eating Questionnaire [Rep(eat)-Q] (Conceição et al., 2017) is a 12-item  
154 self-report questionnaire designed to assess the frequency of attitudinal and behavioural  
155 features of grazing over the past month (e.g., “Ate “on and off” all day without planning it” or

156 “*Could not resist going back to snacking on food even if you were trying to resist doing so*”).  
157 Responses are rated on a 7-point Likert-scale ranging from 0 (never) to 6 (every day), with  
158 higher scores indicating greater frequency. The total score and two subscales, Compulsive  
159 Grazing (CG) and Repetitive Eating (RE), are calculated by averaging the scale items. The  
160 lead author on the development study was contacted who provided supplemental materials  
161 and the items ( Conceicao, Simões, Machado, & Mitchell, 2015; Conceicao, Mitchell, Engel,  
162 et al., 2014). The Rep(eat)-Q was forward and backward-translated in an iterative process by a  
163 bilingual research team of native English and Norwegian speakers (as has been described  
164 previously, see Reas, Bang, Øverås, Lask, & Rø, 2011).

165 The Eating Disorder Diagnosis Scale (EDDS)- DSM-5 Version (Stice, Telch, & Rizvi,  
166 2000) is a brief self-report measure to assess diagnostic features of EDs. The EDDS (Stice et  
167 al., 2000) has shown good reliability and convergent validity, predictive validity, and criterion  
168 validity against diagnoses derived from clinical interviews (Stice, Fisher, & Martinez, 2004;  
169 Sysko et al., 2015). For the present study, the symptom composite score was calculated by  
170 standardizing all items (to control for the effects of the different response formats) and then  
171 summing across all items (except the height and birth control pill items) to provide an overall  
172 indicator of ED pathology. In addition, we used the binge eating frequency item (item 6) (i.e.,  
173 “*how many times per month on average over the past 3 months have you eaten an unusually*  
174 *large amount of food and experienced a loss of control*”). Internal consistency of the EDDS  
175 symptom composite in the present study was .89.

176 The Yale Food Addiction Scale (YFAS) version 2.0 (Gearhardt, Corbin, & Brownell,  
177 2016) is a self-report measure operationalizing addictive-like eating behaviour according to  
178 the DSM-5 diagnostic criteria for substance- use disorders. The YFAS 2.0 demonstrated good  
179 internal consistency, as well as convergent, discriminant, and incremental validity (Gearhardt

180 et al., 2016). We used the continuous symptom count that reflects the number of diagnostic  
181 criteria met (0-11). Internal consistency of the YFAS in the present study was .86.

182 Students volunteered for participation via a university platform for online recruitment  
183 and data collection. All data collected were anonymised and students received course credit  
184 for participation. This study was approved by the Department of Psychology's Institutional  
185 Review Board and Ethics Committee at University of Oslo.

### 186 Statistical Analyses

187 The two factor model described by Conceição et al. (2017) was tested using  
188 confirmatory factor analysis (CFA) with the maximum likelihood discrepancy method.  
189 Confirmatory factor analysis (CFA) is the method of choice to test a priori hypotheses about  
190 latent variables and to evaluate factor invariance across time and groups. Fit indexes indicate  
191 a good fit when:  $\chi^2$  (CMIN) is non-significant ( $p > 0.05$ ); HOELTER.05  $> 200$ ; Standardized  
192 Root Mean Square Residual (SRMR)  $< 0.08$ ; Root Mean Square Error of Approximation  
193 (RMSEA)  $< 0.05$ ; PCLOSE  $> 0.05$ ; Comparative Fit Index (CFI)  $> 0.95$ ; Normed Fit Index  
194 (NFI)  $> 0.95$ ; NNFI (TLI)  $> 0.991$ ; Akaike Information Criterion (AIC) (smaller values  
195 indicate a better model fit) (Hooper, Coughlan, & Mullen, 2008). CFA was conducted using  
196 IBM® SPSS® Amos™ 20.0. Five variables had missing values, all less than 5% missing,  
197 which were replaced with the median values given the variables assess on an ordinal scale.  
198 Modification indexes were examined to determine the co-variance between errors that  
199 improved the model fit. Internal consistency of the Rep(eat)-Q was assessed by Cronbach's  $\alpha$   
200 coefficients. Due to non-normality in the data, Spearman's rank order coefficients were used  
201 to test correlations between the Rep(eat)-Q and self-reported BMI ( $\text{kg/m}^2$ ), age, binge eating  
202 frequency, YFAS 2.0 symptom count, and EDDS composite. Effect sizes were interpreted  
203 according to Cohen's criteria of .1 = small effect, .3 = medium effect, and .5 = large effect



204 (Cohen, 1988; Miles & Shevlin, 2011). These data analyses were performed using IBM®  
 205 SPSS® Statistics 23.0.

## 206 **Results**

207 The Kaiser-Meyer-Olkin measure of sampling adequacy was .905 and the Bartlett's  
 208 test of sphericity was significant ( $p < .001$ ), supporting the appropriateness of factor analysis.  
 209 Confirmatory factor analysis revealed good fit indexes for the 2-factor model to support the  
 210 two subscales: RMSEA = 0.073, CFI = 0.970 and TLI = 0.947; CMIN= 85.982; DF = 43;  $p =$   
 211 0.000; CMIN/df = 2.00; HOELTER.05 = 132; SRMR = 0.048; PCLOSE = 0.050; NFI =  
 212 0.942; AIC = 155.982. Figure 1 depicts the standardized coefficients as well as the  
 213 covariances between errors that improved model fit. Internal consistency was adequate at  
 214 .923, .906, and .864 for the total, CG and RE subscales, respectively. Table 1 summarizes the  
 215 correlations among the Rep(eat)-Q scores, BMI, age, and measures of ED pathology. The  
 216 average total Rep(eat)-Q score for the entire sample was 0.98 (SD: 0.89), and females scored  
 217 significantly higher than males on the total and CG and RE subscales, respectively [Ms = 1.06  
 218 (.95) vs 0.65 (.57),  $t(189) 3.42, p < .001$ ; Ms = 1.16 (.98) vs 0.77 (0.70),  $t(189) 2.88, p =$   
 219 .005, and Ms = 0.95 (1.05) versus 0.53 (.66),  $t(189) 3.09, p = .003$ ].

220 The bivariate correlation between CG and RE for the total sample was  $r_{sp} = 0.69, p <$   
 221 .001. As illustrated in Table 1, age and BMI did not correlate significantly with the Rep(eat)-  
 222 Q, but moderate to strong correlations ( $r_{sp} \geq 0.48$  to 0.61) were found between the CG  
 223 subscale and eating-related measures, whereas weaker, yet significant correlations ( $r_{sp} \geq 0.37$   
 224 - 0.45) were found for the RE subscale. The differential pattern of correlations for CG versus  
 225 RE mirrored the total sample when split by gender, except for weaker, yet significant,  
 226 correlations among males for CG ( $r_{sp} \geq 0.28$  to 0.49), while all correlations between RE and  
 227 eating measures became non-significant ( $r_{sp} \geq 0.11$  to 0.22). For females, the corresponding  
 228 results indicated moderate to strong correlations ( $r_{sp} \geq 0.51$  to 0.63) between the CG subscale

229 and eating-related measures, and moderate correlations for RE ( $r_{sp} \geq 0.42 - 0.51$ ; data not  
230 shown).

### 231 **Discussion**

232 This study supported the utility and validity of the Repetitive Eating Questionnaire  
233 [Rep(eat)-Q] as a brief and easily administered 12-item measure of two subtypes of grazing  
234 behaviour, namely, compulsive grazing (marked by loss of control) and repetitive eating  
235 (eating in a distracted or mindless way). This study represents the first to provide a validation  
236 and replication of the proposed factor structure of the Rep(eat)-Q aside from the initial  
237 development study.

238 Our overall total Rep(eat)-Q score of 0.98 (*SD* 0.89) was lower than previously  
239 reported (i.e.,  $M = 1.59$ , *SD* 1.22) (Conceição et al., 2017), yet this appears to attribute to  
240 disproportionately lower scores in our male participants. Consistent with the subsample of  
241 community participants in the original study (Conceição et al., 2017), women reported a  
242 higher frequency of total and compulsive grazing than men, and we also found higher levels  
243 of repetitive eating in women than men. The observed pattern of associations, in which  
244 stronger associations were found between compulsive grazing and measure of ED pathology  
245 than non-compulsive grazing, was similar for both genders. This invariance indicates that the  
246 proposed differentiating scheme, in which compulsive grazing (marked by a loss of control) is  
247 linked to higher levels of overall eating disorder pathology than non-compulsive grazing,  
248 applies to both genders. This is also in line with recent evidence for the conceptualization of  
249 grazing (in its two subtypes) on a continuum of LOC eating and in the spectrum of ED  
250 psychopathology, in which non-compulsive grazing is associated with the lowest degree of  
251 LOC/psychopathology and compulsive with the highest (Conceicao et al., 2018). In contrast  
252 to the initial development study, however, we found no significant associations between the  
253 frequency of grazing and higher BMI or younger age. It is possible our lack of findings for

254 age or BMI may attribute to our relatively homogenous and smaller sample, whereas the  
255 original study sampled over 1000 participants from different recruitment channels. We note,  
256 however, the correlations originally reported by Conceição et al. (2017) for age and BMI were  
257 quite weak, despite reaching statistical significance, (i.e.,  $r_{sp} = -0.23$  for age and 0.087 for  
258 BMI).

259 Several limitations of the present investigation are noteworthy. This study utilized a  
260 convenience sample of normal weight, educated young adults recruited from a university  
261 setting. Thus, the homogenous sample characteristics place limits on the generalizability of  
262 these findings. Temporal stability was not established in this study, although the original  
263 development study reported good test-retest reliability of the Rep(eat)-Q (0.82). A more  
264 comprehensive assessment battery would have enabled a broader investigation of the  
265 relationship between grazing and other important aspects of mental health (e.g., quality of life,  
266 depression, anxiety) or aspects of eating-related pathology. Longitudinal studies are required  
267 to investigate the prospective validity of grazing as a potential risk behaviour or precursor for  
268 the emergence of other clinically significant disordered eating problems or weight fluctuations  
269 over time.

270 In conclusion, these data provide further confidence in the originally proposed two-  
271 factor model of the Re(peat)-Q and provide support for the utility of this brief assessment to  
272 measure grazing behaviour in non-treatment seeking, community samples of young adults.  
273 This study offers further evidence of two distinct subtypes of grazing behaviour as delineated  
274 by the presence or absence of loss of control, namely, compulsive (marked by loss of control)  
275 and repetitive eating (eating in a distracted or mindless way).

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Table 1. Correlations among Rep(eat)-Q scales, age, BMI, overall ED pathology, binge eating frequency, and YFAS scores (N=191)

	Rep(eat)-Total <sup>1</sup>	Rep(eat)-CG	Rep(eat)- RE
Age	-0.034	0.020	-0.085
BMI	0.057	0.085	0.019
Overall ED pathology <sup>2</sup>	<b>0.575**</b>	<b>0.612**</b>	<b>0.454**</b>
Binge eating frequency <sup>3</sup>	<b>0.462**</b>	<b>0.476**</b>	0.366**
YFAS symptom count <sup>4</sup>	<b>0.489**</b>	<b>0.534**</b>	0.385**

Note. Rep(eat)-CG = Compulsive grazing subscale; Rep(eat)-RE = Repetitive eating subscale; BMI = Body Mass Index; ED = Eating Disorder; YFAS = Yale Food Addiction Scale; <sup>1</sup> = The total mean score from the 12-item Rep(eat)-Q; <sup>2</sup> = The Eating Disorder Diagnostic Scale (EDDS) symptom composite score; <sup>3</sup> = EDDS item 6 (“how many times per month on average over the past 3 months have you eaten an unusually large amount of food and experienced a loss of control”) <sup>4</sup> = The continuous symptom count from the YFAS. Spearman’s rho correlations \**p* < .05, \*\**p* < .01; Correlations in **bold** indicate coefficients  $\geq 0.40$ .