

# BMJ Open Postoperative wound dehiscence after laparotomy: a useful healthcare quality indicator? A cohort study based on Norwegian hospital administrative data

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## ABSTRACT

**Objectives** Postoperative wound dehiscence (PWD) is a serious complication to laparotomy, leading to higher mortality, readmissions and cost. The aims of the present study are to investigate whether risk adjusted PWD rates could reliably differentiate between Norwegian hospitals, and whether PWD rates were associated with hospital characteristics such as hospital type and laparotomy volume.

**Design** Observational study using patient administrative data from all Norwegian hospitals, obtained from the Norwegian Patient Registry, for the period 2011–2015, and linked using the unique person identification number.

**Participants** All patients undergoing laparotomy, aged at least 15 years, with length of stay at least 2 days and no diagnosis code for immunocompromised state or relating to pregnancy, childbirth and puerperium. The final data set comprised 66 925 patients with 78 086 laparotomy episodes from 47 hospitals.

**Outcomes** The outcome was wound dehiscence, identified by the presence of a wound reclosure code, risk adjusted for patient characteristics and operation type.

**Results** The final data set comprised 1477 wound dehiscences. Crude PWD rates varied from 0% to 5.1% among hospitals, with an overall rate of 1.89%. Three hospitals with statistically significantly higher PWD than average were identified, after case mix adjustment and correction for multiple comparisons. Hospital volume was not associated with PWD rate, except that hospitals with very few laparotomies had lower PWD rates.

**Conclusions** Among Norwegian hospitals, there is considerable variation in PWD rate that cannot be explained by operation type, age or comorbidity. This warrants further investigation into possible causes, such as surgical technique, perioperative procedures or handling of complications. The risk adjusted PWD rate after laparotomy is a candidate quality indicator for Norwegian hospitals.

## INTRODUCTION

The past decades have seen a major growth in initiatives for measuring, monitoring and improving the quality of healthcare services. Quality indicators are regularly published

## Strengths and limitations of this study

- Includes all laparotomies performed in the nation over a 5-year period, with patients followed across hospitals.
- Extends previous studies to a new health system and a new coding system.
- The statistical analysis uses methods for low event rates, avoiding asymptotic approximation.
- Results may be subject to coding inaccuracy and incompleteness, as well as selection effects.
- There were no data for surgical technique, nor for some clinical factors known to be relevant.

in many healthcare systems. Performance of healthcare systems is also compared across nations, for instance in the Organisation for Economic Co-operation and Development (OECD) Health Care Quality and Outcomes initiative, which Norway is a part of.<sup>1,2</sup> Norway has a national quality indicator system for monitoring and comparing hospital performance, however, not all areas of hospital performance are covered by existing national quality indicators. While there are quality indicators for outcomes such as mortality and process measures such as waiting times, complications following hospital care is less explored, which is especially relevant following surgical procedures.

Postoperative wound dehiscence (PWD) rates after open abdominal surgery (laparotomies) was introduced as a patient safety indicator in the United States and later as a quality indicator by OECD.<sup>3–6</sup> Norway reported the second highest numbers for 2014–2015, with a PWD rate of 1.02%. The overall range was 0.05% to 1.05%.<sup>5</sup> Neighbouring Sweden, with comparable population health and healthcare, reported 0.30%. Moreover, a recent study comparing adverse events in Norway and Sweden found significantly

higher adverse event rates of surgical complications in Norwegian hospitals, compared with Swedish hospitals.<sup>7</sup>

PWD is a serious complication that leads to higher mortality rates, higher implicit, explicit and social costs as well as increased readmission rates.<sup>8,9</sup> The PWD rate has been studied elsewhere as a quality indicator for hospitals, and found to have a high positive predictive value.<sup>10,11</sup> It is useful as a quality indicator, since several of the risk factors are modifiable and within control of the hospital and surgical team. There are few events per hospital, making it challenging to identify outlier hospitals for quality improvement because of the high statistical uncertainty.

Previous research has identified a number of risk factors for PWD. Examples of such factors are: (I) patient related variables and comorbidities: smoking,<sup>12</sup> obesity,<sup>13</sup> chronic pulmonary disease, renal insufficiency or diabetes<sup>14</sup> and use of immunosuppressive agents<sup>15,16</sup>; (II) procedure related factors: operation type,<sup>9,17</sup> type of incision and closure<sup>18–20</sup> and length of operation time<sup>21</sup>; (III) postoperative parameters: clean wound classification,<sup>21</sup> coughing<sup>9</sup> and wound infection<sup>9,14</sup>; (IV) operative scenario: for example, qualifications of the surgeon<sup>21–23</sup> and of the perioperative team, and whether the surgery is emergent.<sup>9,13</sup>

The objectives of this study are to study the occurrence and variation of PWD after laparotomy at Norwegian hospitals, and the potential usefulness of a PWD indicator for the Norwegian healthcare system, computed from patient administrative data. More specifically we aimed to (1) investigate the possibility to identify hospitals with higher or lower laparotomy PWD rate than average, after appropriate risk adjustment, (2) study the variability of the PWD rate among hospitals, and its relation to hospital type and laparotomy volume.

## MATERIAL AND METHODS

Patient administrative data from all Norwegian hospitals were provided by the Norwegian Patient Registry (NPR) for the period 2011–2015.<sup>24</sup> This comprised individual patient data from all department stays: type of admission (acute or elective), primary and secondary diagnosis codes according to the Norwegian version<sup>25</sup> of ICD-10, surgical and medical procedures, age, gender, date and time of ward admission and discharge. Surgical procedures and operations were coded according to the Norwegian version of the Nordic Medico-Statistical Committee (NOMESCO) Classification of Surgical Procedures.<sup>26</sup> Procedure time and date were not available. It was therefore not possible to exclude reclosures of wounds occurring before or on the same day as laparotomies within the same episode, as requested in the OECD indicator specification. The NPR data files were checked for missing values and inconsistencies between variables, such as date and time of discharge before admission or invalid ICD-10 code. We had no access to clinical data, such as, eg, type of

suture, which would have enabled us to study the causes of the reported dehiscences.

Wound dehiscence was defined as the occurrence of a code for a reclosure operation, that is, a reoperation for wound dehiscence. This excludes superficial dehiscences, as these are usually not resutured, and the code for reclosure operation is restricted to deep wound dehiscences. Laparotomies and wound reclosure operations were identified according to procedure codes. An operation coded with a laparotomy code, signifies an incision into the abdominal wall, through the fascia and with an opening of the abdominal cavity. Laparoscopic and endoscopic procedures were not included. Details of the codes used can be found in the online supplementary file.

All permanent residents in Norway have a personal identification number (PIN), registered in the NPR. NPR prepared an encrypted PIN for all patients with a valid PIN, allowing tracking of patients over time and between hospitals. The data were linked with the National Registry to provide data of death (when applicable), using the PIN.

Ward admissions for each patient, at more than one hospital in case of transfers, were linked into episodes of care when less than 8 hours elapsed from time of discharge to the next ward admission.<sup>27</sup> An episode was regarded as acute if the first admission in the episode was coded as non-elective, as a laparotomy episode if it included any procedure code for laparotomy (reclosures not included), and a reclosure episode if a reclosure code was found. The initial data set consisted of all laparotomy and reclosure episodes. Each reclosure episode was linked to the laparotomy episode immediately preceding or coinciding with it. Reclosure episodes with no preceding laparotomy episode within 30 days, as well as laparotomy or reclosure episodes following a reclosure episode within 30 days, were excluded. Note that the linking of laparotomies and reclosures was not part of the original OECD specification, but is required in order to attribute PWD to hospitals and to enable risk adjustment. Following the OECD specification, laparotomy episodes (and consequently any linked reclosure episodes) were excluded if a diagnosis code for immunocompromised state or relating to pregnancy, childbirth and puerperium was present, if the length of stay was less than 2 days, or if the patient's age was less than 15 years. Hospitals with less than 10 laparotomies over the 5-year period were excluded. The hospitals belonged to one of three types: regional, large with acute function and small with acute function. For details of the diagnosis and operation codes used, see the online supplementary file. For risk adjustment, Charlson comorbidities were determined from previous admissions 3 years prior to, but not including the current episode of care.<sup>27–29</sup> Diagnoses were grouped according to the Clinical Condition Summary system, adapted to the Norwegian version of ICD-10.<sup>30</sup>

## Statistical methods

Risk adjusted probabilities for a laparotomy episode resulting in a reclosure operation were estimated by bias corrected logistic regression.<sup>31</sup> The final model was fit by stepwise

regression with the BIC criterion, allowing for potential two-way interactions.

To identify outlier hospitals, that is, those with high or low risk adjusted PWD probabilities, estimated hospital effects were compared with a reference value, defined as the 25% trimmed mean of the hospital effects on the logistic scale.<sup>32</sup> As some hospitals reported zero reclosures, ordinary maximum likelihood estimates of the model parameters do not exist, due to separation,<sup>33</sup> and the estimated variances of the fitted parameters, based on their asymptotic distribution, become unreliable. The comparison used an exact test based on the Poisson binomial distribution for the number of PWDs per hospital, using the estimated probabilities for each case, together with parametric bootstrapping to account for the estimation uncertainty in the model parameters. Tests for significance were corrected for multiple comparisons using the Guo-Romano method,<sup>34</sup> and outlier status assigned according to the false discovery rate (FDR). An FDR not exceeding 5% was regarded as significant. For sensitivity analysis, two alternative risk adjustment models were tested, with either a four-category grouping of procedures, or with diagnosis categories, instead of the 13-category procedure grouping. In addition, a model with the four Norwegian hospital regions was also estimated.

Hospital volume, modelled by splines,<sup>35</sup> was tested for inclusion in the model. We also performed this test after exclusion of hospitals with zero reclosures.

Finally, the hospital specific effects were modelled as a mixture of two normal distributions. The expectation-maximisation algorithm was used, taking into account the

estimation variances. The mixture model yielded estimates of the quartiles of the hospital ORs and the scaled IQR (normalised by dividing by 1.349, to give the SD in the case of a normal distribution) was computed as a measure of spread among hospitals. Bootstrapping of the mixture model was used to find a 95% CI for the scaled IQR.

### Risk adjustment

The following case-mix variables were included as candidates in the stepwise regression: age, gender, indicators for the individual Charlson comorbidities, number of previous hospital admissions 2 years prior to current admission, and whether the episode was acute or elective. A linear trend in admission year was also included. Age was modelled by natural splines with knots at the median and quartiles.<sup>35</sup> Based on previous studies of risk factors,<sup>9 17</sup> procedures were categorised into 13 types, according to the body system or organ involved. The effects of operation types were normalised to have zero sum on the logistic scale.

For a quality indicator, only characteristics of the patient when entering the hospital, are meaningful risk adjustment variables. No data were available for smoking, obesity or other patient or case characteristics such as nutritional status. There was no information about operation urgency beyond the status of the hospital admission or episode as elective or acute.

### Patient and public involvement

Patients were not involved in the planning, conduct or analysis of this study. The policy of the Norwegian Institute

**Table 1** Descriptive statistics for final data set

	PWD	No PWD
Age, years, median (quartiles)	69 (61–78)	65 (51–75)
Gender, females, n (%)	517 (35)	43 094 (56)
Acute laparotomy episode, n (%)	657 (44)	26 381 (34)
Main diagnosis for reclosure episode coded as PWD, n (%)	45 (3.1)	—
Main diagnosis for reclosure episode coded as deep wound infection, n (%)	274 (19)	—
Hospital type for laparotomy episodes		
Regional, n (%)	545 (37)	28 104 (37)
Large with acute function, n (%)	810 (55)	40 291 (53)
Small with acute function, n (%)	122 (8.3)	8214 (11)
Comorbidities		
Diabetes with complications, n (%)	18 (1.2)	893 (1.2)
Chronic pulmonary disease, n (%)	196 (13)	5147 (6.7)
Renal disease, n (%)	66 (4.5)	2716 (3.5)
30-day mortality (laparotomy episode), %	67 (4.5)	2668 (3.5)
Length of stay laparotomy episode, days, median (quartiles)	19 (11–29)	7.4 (4.4–13)
Reclosure and matched laparotomy in same episode, n (%)	1211 (82)	—
Converted from laparoscopy or endoscopy to laparotomy, n (%)	12 (0.81)	578 (0.75)
Robot assistance in laparotomy, n (%)	3 (0.2)	404 (0.53)

PWD, postoperative wound dehiscence.

**Table 2** Final multivariate logistic model for risk adjustment

Variable	Adjusted ORs (95% CI)
Year of admission	0.93 (0.90 to 0.96)
Age, spline function	
40 (reference)	1.00
50	1.37 (1.25 to 1.49)
60	1.97 (1.65 to 2.36)
70	2.39 (1.97 to 2.90)
Gender	
Female (reference)	1
Male	2.42 (2.16 to 2.72)
Elective laparotomy episode (reference)	1
Acute laparotomy episode	1.36 (1.21 to 1.52)
Chronic pulmonary disease	1.72 (1.47 to 2.01)
Operation type*	
Exploratory laparotomy	2.40 (1.78 to 3.24)
Hernia (diaphragmal)	2.57 (1.37 to 4.81)
Thoracoabdominal aorta	2.08 (0.85 to 5.09)
Gastrointestinal tract	2.04 (1.69 to 2.46)
Liver	1.14 (0.69 to 1.87)
Biliary tract	0.12 (0.05 to 0.28)
Pancreas	0.79 (0.40 to 1.58)
Spleen	1.20 (0.45 to 3.24)
Other digestive system	1.46 (1.03 to 2.07)
Kidney	0.09 (0.03 to 0.28)
Other urinary and male genital organs	0.52 (0.37 to 0.71)
Female genital organs	1.43 (1.06 to 1.92)
Peripheral vascular surgery	1.21 (0.93 to 1.57)
More than one type of surgery†	2.58 (2.12 to 3.15)
Hospital	
Scaled IQR	0.30 (0.23 to 0.34)

\*ORs for operation type is scaled to have geometric mean of one.

†Not counting exploratory laparotomy.

of Public health is to publish hospital quality indicators, when they have been successfully validated.

## RESULTS

The initial data set comprised 96 102 episodes with laparotomy and 1909 with a reclosure operation. After restricting data to reclosures paired with a laparotomy within 30 days, 1580 reclosures remained. After exclusions for pregnancy, childbirth and puerperium or immunocompromised state, age and length of stay, 78 299 laparotomies remained. Lastly, hospitals with less than 10 laparotomies were excluded, yielding a final data set with 66 925 unique patients, 78 086 laparotomies and 1477 reclosures from 47 hospitals. Descriptive statistics for

the dataset are shown in [table 1](#). The operation types are tabulated in the online supplementary file.

From 2011 to 2015, the annual volume of laparotomies decreased somewhat, from 16 730 to 14 419, while the proportion of acute laparotomies remained stable at around 35%.

The overall rate of PWD for the 5-year period was 1.89%. Crude PWD rates varied from 0% to 5.1% among hospitals. After risk adjustment, the range was 0.1%–5.4%. [Table 2](#) shows the ORs of the final logistic regression model. No interactions were included. The model showed good fit according to the modified Hosmer-Lemeshow test<sup>36</sup> ( $p=0.53$ ) and good predictive ability, with an area under the operating characteristic ( $c$ -statistic) of 0.73.

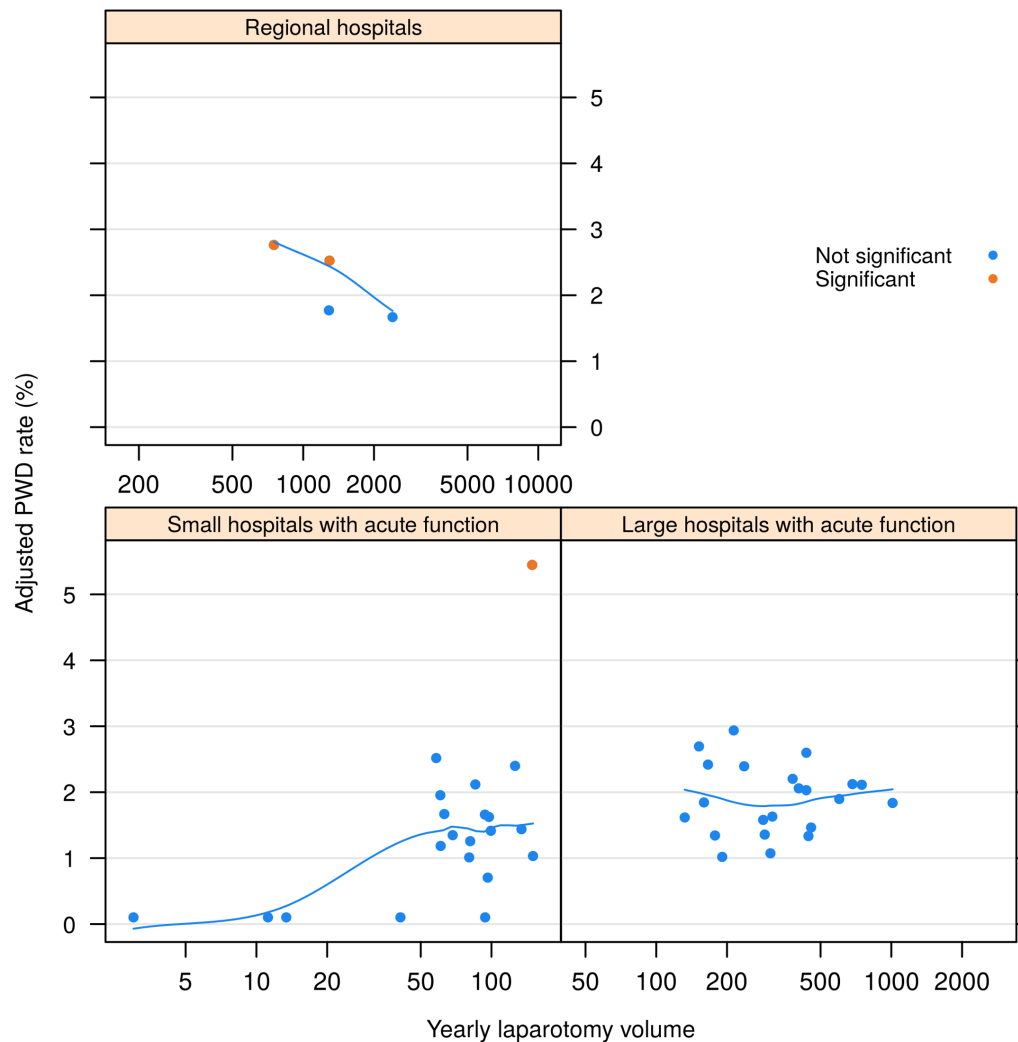
In [figure 1](#), risk-adjusted PWD rates are shown for each hospital, plotted versus laparotomy volume and hospital type.

After significance testing, we identified three hospitals with higher PWD and none with lower PWD than average, when correcting for multiple testing. Without multiple test correction, one additional hospital with high PWD was found to be marginally significant ( $p=0.053$ ).

In the alternative model including volume, the PWD increased with yearly laparotomy volume from a very low level up to 120 laparotomies per year, after which it remained fairly constant, see [figure 1](#). The effect of volume was significant ( $p<0.001$ ), also after exclusion of the four smallest volume hospitals with zero reclosures ( $p=0.008$ ). Hospital type coincided almost completely with a grouping of hospitals by volume, and was therefore not tested separately. There was significant variation among regions ( $p<0.001$ ), with the Northern region having the highest and the South-Eastern region the lowest rates. Details can be found in the online supplementary file. Using diagnosis categories or aggregated operation type as risk adjustment variables resulted in very small changes in risk adjusted PWD rates.

## DISCUSSION

We have studied wound dehiscence after laparotomy, as a quality indicator based on the OECD specification, and found that it discriminated between Norwegian hospitals. The indicator was risk adjusted for differences in age, gender, comorbidity and type of surgery, and showed little sensitivity to changes in the set of risk adjustment variables. The overall PWD rate was 1.89%. After risk adjustment, the hospitals' PWD rate varied between 0.1% and 5.4%. Laparotomy volume and type of hospital had little effect on the PWD rate, except for hospitals with very low volume. Advanced age, male gender, chronic pulmonary disease and emergency laparotomy were all significant risk factors for PWD. There were significant PWD differences according to the organ system targeted. The overall rate of PWD showed a small but statistically significant decline over the observation period 2011–2015. The relatively large variation of PWD rates between hospitals, after



**Figure 1** Risk adjusted PWD rates versus yearly laparotomy volume, by hospital type. Trend curve is obtained by smoothing the scatterplot. Significance testing is adjusted for multiple comparisons. PWD, postoperative wound dehiscence .

correction for patient characteristics and operation type, indicates possible variation in the quality of healthcare among hospitals. This may be due to variation in surgical technique and perioperative care, as well as the handling of postoperative complications, such as wound infection, which is known to be a risk factor for PWD.<sup>17</sup> We found PWD rates well within the range reported in international studies.<sup>9 13 14 17 21 37 38</sup> Also, the risk factors identified are in accordance with previous studies, although limited to administrative data. Laparotomy volume has negligible effect apart from the few hospitals with very low volume. A Japanese study reported a similar conclusion, while volume was found to have effect in US hospitals.<sup>39 40</sup> The effect is likely a result of the types of operations performed at the low-volume hospitals, compared with the other hospitals.

Our study is based on complete data from all Norwegian hospitals performing laparotomies. It was possible to track patients during transfers and reoperations at different hospitals. To the best of our knowledge, no similar study has been performed. NPR, the data source, has been validated for several disease categories with respect to

identification of cases based on diagnoses and/or procedures, and found to have a very high degree of completeness, compared with Norwegian national medical quality registries.<sup>41–44</sup> At the time of writing, the completeness of NPR, after 24 registries have been studied, ranges from 83.5% to 99.8%.<sup>24</sup>

We cannot exclude a residual imbalance in case mix, affecting PWD through eg, smoking or obesity, which are known risk factors. There is regional variation in the prevalence of smoking and obesity in Norway.<sup>45</sup> Obesity is more prevalent in Northern Norway, where PWD rates are somewhat higher. However, in some other areas where obesity is less prevalent, the rates are similar. There is no consistent correspondence between the known variation in smoking among counties and PWD rates. Some surgical procedures are performed only at regional hospitals, and it is therefore possible that selection effects are present. In that case, one would expect larger changes in PWD rates after risk adjustment for operation type, which was not found. One potential source of error in our study is the completeness and correctness of coding in the NPR, particularly the coding of reclosure operations. The

risk adjustment depends on data from previous hospitalisation and may not capture all comorbidities. Moreover, selection effects cannot not be ruled out. Differing policies for operations on patients with known risk factors, for example, obesity or smoking, would likely cause variation in PWD rates. Patients who die before reoperation or are managed by other means will not be registered. We believe that this applies to very few patients and would not influence our results. No attempt was made to identify main operation or operation intent, as this would require a classification effort outside the scope of the present study. No clinical details about surgical technique and patient condition were available. Therefore, the causes of the observed PWD rate variation could not be investigated.

Previous studies have shown that the quality indicator has high positive predictive value, but only moderate sensitivity.<sup>10 46 47</sup> Since we have used specific wound reclosure codes, similar to those used in previous studies, we expect a high positive predictive value in Norway as well. Conceivably, the sensitivity depends on the coding system, in particular the various alternative codes related to complications. Sensitivity in Norway may thus differ from that of other healthcare systems. A recent retrospective medical record study from neighbouring Sweden reports that 86.9% of wound dehiscences were reoperated.<sup>38</sup> Norway has an activity-based system for financing hospitals, which is an incentive to report all reclosure operations.

## CONCLUSIONS

Among Norwegian hospitals, there is a significant variation in PWD rate after laparotomies that cannot be explained by operation type, age, comorbidity or whether the admission was elective or acute. This warrants further investigation into possible causes, such as patient related factors, surgical technique, perioperative procedures or handling of complications, for example, wound infections. Some of these factors are known to be amenable.<sup>20 48</sup> The relatively large between-hospital variation found in the present study is an indication of potential for improvement. The risk adjusted PWD rate after laparotomy is a candidate for use as a quality indicator for Norwegian hospitals, and will make it possible to identify hospitals with apparent quality problems. To achieve sufficient discrimination, however, 5-year data are desirable, making it more difficult to monitor changes in hospital performance resulting from quality improvement efforts. It lies outside the scope of the present study to perform a comprehensive validation of the PWD rate as a quality indicator suitable for public reporting. There are uncertainties and potential biases in the indicator, implying that it must be regarded as a signal for follow-up within hospitals, rather than giving a final verdict of inferior or superior quality. For reporting on surgical quality, several indicators should be used to give a balanced view of the different aspects of quality and patient safety.

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**Contributors** AKL and OT conceived the study. DTK, TMH, OT and SH participated in data preparation. OT, SH and AKL contributed to the analysis. OT helped draft the manuscript. JH was responsible for the statistical analysis and final manuscript. All authors revised and approved the final manuscript.

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